Statistical Methods III: Spring 2013

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 $\mathsf{QP}+\mathsf{constrained}$ inference

Outline







OLS

Scalar notation

$$\underset{\beta}{\operatorname{arg\,min}} \sum (y_i - x^{\top}\beta)^2$$

Matrix notation

$$rgmin_{eta}(Y-Xeta)^{ op}(Y-Xeta)$$

Matrix notation expansion

$$\underset{\beta}{\arg\min} \beta^{\top} X^{\top} X \beta - 2 Y^{\top} X \beta + Y^{\top} Y$$

• No constraints on β

Reviewing (and rewriting) OLS

Matrix notation expansion

$$\underset{\beta}{\operatorname{arg\,min\,}\beta^{\top}X^{\top}X\beta} - 2Y^{\top}X\beta + Y^{\top}Y$$

multiply by 1/2, and drop YY (why allowed?)

$$\operatorname*{arg\,min}_{\beta} \frac{1}{2} \beta^{\top} X^{\top} X \beta - Y^{\top} X \beta$$

If we write in form of,

$$\underset{\beta}{\operatorname{arg\,min}} d^{\top}\beta + \frac{1}{2}\beta W\beta$$

then

OLS - with restrictions

OLS solution places no restriction on β . The optimization problem

$$\operatorname*{arg\,min}_{eta} rac{1}{2} eta^{ op} X^{ op} Xeta - Y^{ op} Xeta$$

restricted inequality constraints on β , for example,

$$\beta \ge \mathbf{0}$$

can be solved with "quadratic programing"

Quadratic programming

Definition (Quadratic programming problem)

Let $\beta \in \mathbb{R}^n$, *W* be symmetric $n \times n$ matrix,

$$\operatorname*{arg\,min}_{\beta} {\pmb{\sigma}}^{\top} {\pmb{\beta}} + \frac{1}{2} {\pmb{\beta}}^{\top} {\pmb{W}} {\pmb{\beta}}$$

subject to

$$R_1^{\top}\beta \ge b_1$$

 $R_2^\top \beta = b_2$

Notes

• *d*, *W*, *R*, and *b* are fixed for a given optimization

• β is unknown









Constrained estimation: dimensions and distance Constrained estimates $\tilde{\Delta}$, subject to $\Delta > 0$

$$\tilde{\Delta} = \begin{cases} (\hat{\Delta}_1, \hat{\Delta}_2) & \text{if } \hat{\Delta}_1 \geq 0 \text{ and } \hat{\Delta}_2 \geq 0 \\ (\hat{\Delta}_1, 0) & \text{if } \hat{\Delta}_1 \geq 0 \text{ and } \hat{\Delta}_2 < 0 \\ (0, \hat{\Delta}_2) & \text{if } \hat{\Delta}_1 < 0 \text{ and } \hat{\Delta}_2 \geq 0 \\ (0, 0) & \text{if } \hat{\Delta}_1 < 0 \text{ and } \hat{\Delta}_2 < 0 \end{cases}$$

How many free parameters? 2 (S1), 1 (S2, S4), or 0 (S3). What is probability of being "far" from H_0 ? By quadrant:

$$egin{aligned} & \mathcal{P}(ilde{\Delta}_1^2 + ilde{\Delta}_2^2 < m{c}' \mid \hat{\Delta} \in m{S}_1) = \mathcal{P}(\chi_2^2 < m{c}') \ & \mathcal{P}(ilde{\Delta}_1^2 + 0 < m{c}' \mid \hat{\Delta} \in m{S}_2) = \mathcal{P}(\chi_1^2 < m{c}') \ & \mathcal{P}(0 + ilde{\Delta}_2^2 < m{c}' \mid \hat{\Delta} \in m{S}_4) = \mathcal{P}(\chi_1^2 < m{c}') \ & \mathcal{P}(0 + 0 < m{c}' \mid \hat{\Delta} \in m{S}_3) = 1 \end{aligned}$$

Constrained estimation: dimensions and distance

In this simple example, each of quadrants is equally likely under H_0

$$egin{aligned} & \mathcal{P}(ilde{\Delta}_1^2 + ilde{\Delta}_2^2 < c' \mid \hat{\Delta} \in S_1) = \mathcal{P}(\chi_2^2 < c') \ & \mathcal{P}(ilde{\Delta}_1^2 + 0 < c' \mid \hat{\Delta} \in S_2) = \mathcal{P}(\chi_1^2 < c') \ & \mathcal{P}(0 + ilde{\Delta}_2^2 < c' \mid \hat{\Delta} \in S_4) = \mathcal{P}(\chi_1^2 < c') \ & \mathcal{P}(0 + 0 < c' \mid \hat{\Delta} \in S_3) = 1 \end{aligned}$$

So for given α , solve for c'

$$P(\bar{\chi}^2 < c') = 1/4P(\chi_2^2 < c') + 1/2P(\chi_1^2 < c') + 1/4 = 1 - \alpha$$

Distribution of hypothesis tests

- Equality restriction versus unconstrained
 - Fixed difference in dimensionality of models

 $\Pr(-2[L(\tilde{\theta}) - L(\hat{\theta})] > c) = \Pr(\chi_r^2 > c)$

But number of free – restricted parms is stochastic if

- Equality restrictions versus inequality restriction
- Inequality restricted vs unconstrained
- Inequality restricted versus additional inequalities

$$Pr(-2[L(\tilde{\theta}) - L(\hat{\theta})] > c') = Pr(\bar{\chi}^2 > c')$$
$$= \sum_{k=1}^{K} w_k Pr(\chi_k^2 > c')$$

where w_k is the probability of having a difference of k degrees of freedom between models

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Overlap of sets: $H_0 vs_{\Delta_2} H_{\emptyset}$ and $H_0 vs H_{\nearrow}$



Sequence of means



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Best fitting mono inc.

Isotonic regression isoreg(x = x, y = y)



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Diff in adjacent mean, unconstrained to constrained



Sequence of means



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Best fitting mono inc.

Isotonic regression isoreg(x = x, y = y)



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Diff in adjacent mean, unconstrained to constrained



Lots of draws of unconstrained diff means



Constrained v mono inc diff



Constrained v mono inc diff

In simulation,

- could you determine w, probability of model dimensionality
- what is the distribution of the fits in constrained?

Transformation via cholesky



Transformation via cholesky

```
Y <- matrix(rnorm(3*n),ncol=3)
d <- t(apply(Y,1,diff))
z <- t(apply(Y,1,function(x) isoreg(x)$yf ))
d2<- t(apply(z,1,diff))</pre>
```

```
(V <- cov(d))
Vi<- solve(V)
A <- t(chol(Vi))
Ai<- solve(A)
d3 <- d %*% A</pre>
```

... now iid bivariate normal



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Transformed optimization problem...

```
d <- t(apply(Y,1,diff))
d2<- t(apply(z,1,diff))
(V <- cov(d))
Vi<- solve(V)
A <- t(chol(Vi))
Ai<- solve(A)
d3 <- d %*% A</pre>
```

```
X <- d[i,]
d4 <- qp2(X)
```

Constrained fit on transformed means



Constrained v mono inc diff



Transformed vs untransformed

In simulation,

- R and V determine cone/constraint
- note: perpendicular: closest fitting point compare with untransformed
- could you determine *w*, probability of model dimensionality (any different from untransformed?)
- what is the distribution of the fits in constrained?